

# Effectiveness of vaccination against SARS-CoV-2 Omicron variant infection, symptomatic disease, and hospitalization: a systematic review and meta-analysis

Angela Meggiolaro, Monica Sane Schepisi, Sara Farina, Carolina Castagna, Alessia Mammone, Andrea Siddu, Paola Stefanelli, Stefania Boccia & Giovanni Rezza

To cite this article: Angela Meggiolaro, Monica Sane Schepisi, Sara Farina, Carolina Castagna, Alessia Mammone, Andrea Siddu, Paola Stefanelli, Stefania Boccia & Giovanni Rezza (2022) Effectiveness of vaccination against SARS-CoV-2 Omicron variant infection, symptomatic disease, and hospitalization: a systematic review and meta-analysis, Expert Review of Vaccines, 21:12, 1831-1841, DOI: [10.1080/14760584.2022.2130773](https://doi.org/10.1080/14760584.2022.2130773)

To link to this article: <https://doi.org/10.1080/14760584.2022.2130773>



© 2022 Italian Ministry of Health. Published by Informa UK Limited, trading as Taylor & Francis Group.



[View supplementary material](#)



Published online: 14 Oct 2022.



[Submit your article to this journal](#)



Article views: 1301



[View related articles](#)



[View Crossmark data](#)



Citing articles: 1 [View citing articles](#)

# Effectiveness of vaccination against SARS-CoV-2 Omicron variant infection, symptomatic disease, and hospitalization: a systematic review and meta-analysis

Angela Meggiolaro<sup>a</sup>, Monica Sane Schepisi<sup>a</sup>, Sara Farina<sup>b</sup>, Carolina Castagna<sup>b</sup>, Alessia Mammone<sup>a</sup>, Andrea Siddu<sup>a</sup>, Paola Stefanelli<sup>c</sup>, Stefania Boccia<sup>b,d</sup> and Giovanni Rezza<sup>a</sup>

<sup>a</sup>General Directorate for Health Prevention, Italian Ministry of Health, Viale Ribotta 5, 00144 Rome, Italy; <sup>b</sup>Department of Life Sciences and Public Health, Section of Hygiene, Università Cattolica del Sacro Cuore, 00168 Rome, Italy; <sup>c</sup>Department of Infectious Diseases, Istituto Superiore di Sanità, Viale Regina Elena 299, 00161 Rome, Italy; <sup>d</sup>Department of Woman and Child Health and Public Health-Public Health Area, Fondazione Policlinico Universitario A. Gemelli IRCCS, 00168 Rome, Italy

## ABSTRACT

**Background:** This meta-analysis aims to assess the effectiveness of the current Sars-Cov2 vaccine regimens against Omicron infection. A secondary endpoint aims to investigate the waning effectiveness of primary vaccination against symptomatic infection and related hospitalization.

**Research design and methods:** The systematic review started on 1 December 2021 and was concluded on 1 March 2022. Random-effects frequentist meta-analyses and multiple meta-regressions were performed.

**Results:** In total, 15 studies are included in the quantitative synthesis. According to the meta-analysis results, the overall risk of Sars-Cov2 infection in vaccinated individuals is on average 31·5% lower than the infection risk in unvaccinated while vaccinated with one booster dose have a 70·4% risk reduction of Omicron infection compared to unvaccinated. In particular, one booster dose significantly decreases by 69% the risk of symptomatic Omicron infection with respect to unvaccinated. Six months after the primary vaccination, the average risk reduction declines to 22% against symptomatic infection and to 55% against hospitalization.

**Conclusions:** Primary vaccination does not provide sufficient protection against symptomatic Omicron infection. Although the effectiveness of the primary vaccination against hospitalization due to Omicron remains significantly above 50% after 3 months, it dramatically fades after 6 months.

## ARTICLE HISTORY

Received 5 July 2022  
Revised 1 September 2022  
Accepted 27 September 2022

## KEYWORDS

Booster dose; Sars-Cov2 vaccine; effectiveness; omicron VOC Sars-Cov2 vaccine; symptomatic omicron infection; risk of hospitalization; waning immunity

## 1. Introduction

On 26 November 2021, the WHO designated the variant B.1.1.529 (named Omicron) as a variant of concern. The global epidemiology of SARS-CoV-2 has been characterized by the rapid spreading of the Omicron variant (B.1.1.529) and Omicron has become the dominant variant circulating globally ever since [1]. To date, Omicron encompasses several sub-lineages, the most common ones being BA.1, BA.1.1, and BA.2.

The SARS-CoV-2 Omicron variant contains several important mutations on the spike protein, potentially leading to deleterious consequences. The increased transmissibility of Omicron is determined by a combination of i) intrinsic biological properties that make the virus more infectious than previous lineages (e.g. ACE2 receptor-binding efficiency or viral replication efficiency) [2,3], and ii) immune escape properties resulting in more outbreaks among vaccinated or more reinfections among recovered individuals [4,5].

Regarding the clinical severity, a less severe onset, lower hospital admission rates and/or shorter length of hospital stay,

as well as declining case fatality rates have been extensively documented by the scientific literature [6–9].

COVID-19 vaccines licensed in the EU have proven highly effective in preventing SARS-CoV-2 infections [10–14]; however, several in vitro studies suggest a reduction in neutralizing titers against Omicron in individuals who have received vaccination with two or three doses and in those who have had prior SARS-CoV-2 infection [15–17]. Clinical studies have suggested that the levels of antibodies after BNT162b2, mRNA-1273 and Ad26.COV2.S vaccines could last for at least 6 months and decrease over time thereafter [18–20]. Nonetheless, recent findings on cross-neutralizing immunity against Omicron among individuals that received a third dose of mRNA vaccine suggest that the current vaccine regimens may still overcome evasion of humoral immunity [21].

Omicron variant's higher transmissibility combined with an increased risk of infection among vaccinated individuals has prompted health authorities to consider the introduction of a booster dose [22]. Therefore, estimating whether and how

Sars-Cov2 primary vaccination effectiveness fades over time is essential to pinpoint the optimal timing for the booster dose.

The objective of this meta-analysis is twofold: first, to assess Sars-Cov2 vaccine effectiveness against infection, symptomatic disease, and hospitalization due to laboratory-confirmed SARS-CoV-2 Omicron variant. Second, to investigate the waning effectiveness of the primary course vaccination against Omicron over time.

## 2. Methods

### 2.1. Search strategy and selection criteria

This systematic review, with meta-analysis, is based on a web search updated weekly until 1 March 2022 (**Table S1, Supplementary material**). The sources of information essentially consist of three web engines, including early-stage research platforms (i.e. WHO COVID-19 DATABASE, PubMed, medRxiv + bioRxiv), all relevant web resources reporting living data on vaccine effectiveness (i.e. <https://view-hub.org/covid-19/> and <https://covid-nma.com/>), electronic databases, and gray literature. Reviews and their references are examined for inclusion. No country, language, study design restrictions are applied.

All the relevant records are screened by title and abstract. Potentially relevant publications undergo full-text examination and disagreements on eligibility are solved through discussion by all the authors. The full texts suitable for the quantitative synthesis are collected in an excel database for data extraction. The items for data extraction are predefined and agreed upon by all authors. The systematic review and meta-analyses are performed in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) 2020 Statement guidelines [23] (**PRISMA checklist: Supplement 1**) This study is registered with PROSPERO, CRD42021240143. (<https://www.crd.york.ac.uk/PROSPERO/>)

### 2.2. Data extraction

Data extracted by at least three out of five independent investigators are collected in Excel tables. The information drawn up from each full text include the following:

1. General characteristics of the study: design, year of publication, country, mean age of the sample, follow-up, risk of bias;

2. Exposure: data are stratified according to the Sars-Cov2 vaccination course; hence, two main groups are acknowledged, corresponding to primary vaccination and one additional booster-dose recipients. Within each subgroup, the vaccination course is classified according to the vaccine type (ChAdOx1 nCoV-19, Ad26.COV2.S, BNT162b2, or mRNA-1273 vaccine). This meta-analysis does not include immunization regimens created with inactivated vaccines such as CoronaVac. Heterologous primary schedules are included. All SARS-CoV-2 vaccine recipients are considered as *exposed*, while unvaccinated are considered as *unexposed*.

3. Outcome: cases are defined as being due to the Omicron variant, based on S target–negative results on PCR or whole-genome sequencing. Regardless of the vaccine course

undertaken, cases occurred within 14 days after the primary vaccination or within 1 week from the booster administration are not included. Omicron cases are classified by clinical severity into any Sars-Cov2 infection excluding hospitalization, symptomatic disease, and hospitalization due to COVID-19 disease.

4. Risk of bias: ROBINS-I (risk of bias in non-randomized studies of interventions) is applied to assess risk of bias. The tool classifies the risk into 'low,' 'moderate,' and 'serious' [24].

### 2.3. Endpoints

The primary endpoint aims to assess the overall effectiveness of the current Sars-Cov2 vaccination regimens against Omicron. The study results are stratified by clinical severity and reported at the maximum follow-up.

The secondary endpoint attempts to measure the waning effectiveness of the primary vaccination at consecutive time intervals. In particular, VE is assessed in intervals of 3, 6, and more than 6 months after the last dose.

The point estimates of the effect size, as measured by Log odds ratios (Log ORs) and 95% confidence interval (95% CI), are computed through meta-analysis and converted to ORs by exponentiation. VE is quantified as the risk reduction of any infection event, expressed as a percentage, compared to the unvaccinated group.

### 2.4. Statistical analysis

A random-effects (RE) model employing inverse variance method (IV) is fitted to the data. The amount of heterogeneity (i.e.  $\tau^2$ ) is estimated using the restricted maximum-likelihood estimator [25]. In addition to the estimate of  $\tau^2$ , the QQ-test for heterogeneity [26] and the  $I^2$  statistic [27] are reported. Studentized residuals and Cook's distances are used to examine whether studies may be outliers and/or influential in the context of the model. The normality assumption is evaluated via QQ normal plot [28].

The publication bias is evaluated through a funnel plot and tested via regression test (weighted regression with multiplicative dispersion). The rank correlation test [29,30] and the regression test [30], using the standard error of the observed outcomes as predictor, are used to check for funnel plot asymmetry.

Regarding the primary vaccination waning effectiveness, the subgroup meta-analyses include the stratification by time intervals since the last dose uptake for symptomatic Covid-19 risk and hospitalization risk due to Sars-Cov2 infection. Studies providing vaccine effectiveness estimates at discrete time intervals after the primary vaccination course, which met the predefined screening criteria, underwent further meta-analysis and meta-regression. In order to test for subgroup differences, both a mixed-effect meta-regression model assuming a common  $\tau^2$  value within the subgroups and a three-level meta-regression model, allowing for different  $\tau^2$  values across subgroups, are fitted.

Finally, in order to examine whether one or multiple moderator variables are able to account for the heterogeneity (or part of it), multiple meta-regressions are performed under the

mixed-effects model for both continuous and nominal study level covariates [25]. The analysis is carried out using R (version 4.0.5).

### 2.5. Role of the funding source

There was no funding source for this study.

## 3. Results

The web search provided 502 unduplicated records (Figure 1). In total, 15 studies and 55 observations are included in the quantitative synthesis concerning the overall Sars-Cov2 vaccine effectiveness against Omicron VOC. All of them have a test-negative case-control design except one cohort study [31]. (Table S2, Supplementary material). The majority of the studies are carried out in the US and the UK (59%), the sample age is on average 45 years, while the induction period for immunization appears slightly shorter for studies analyzing

the effectiveness of the booster dose compared to those investigating the primary course vaccination (on average 11 and 16 days since administration, respectively). The 75% of the observations concern the mRNA vaccine effectiveness, while the 13% involve heterologous vaccine regimens. The booster dose is administered on average 6 months after the primary course (Table S1, Supplementary material 2). The majority of the selected studies involves the general population, while Gray et al. report results on HCWs (32)[Gray] and Spensley et al. on patients affected by end-stage kidney disease receiving in hospital hemodialysis [31]. All the studies examine the VE of mRNA BNT162b2 vaccine except Tseng et al. and Gray et al. which investigate mRNA-1273 and aAd26.COV.2 effectiveness, respectively [32,33].

Five studies report data on the waning effectiveness of the primary vaccination against symptomatic Omicron infection [34–38]. A considerable effectiveness rebound after mRNA booster dose is shown by four studies [34–36,38]. VE against hospitalization caused by Omicron is analyzed by seven studies [32,33,36,39–42].

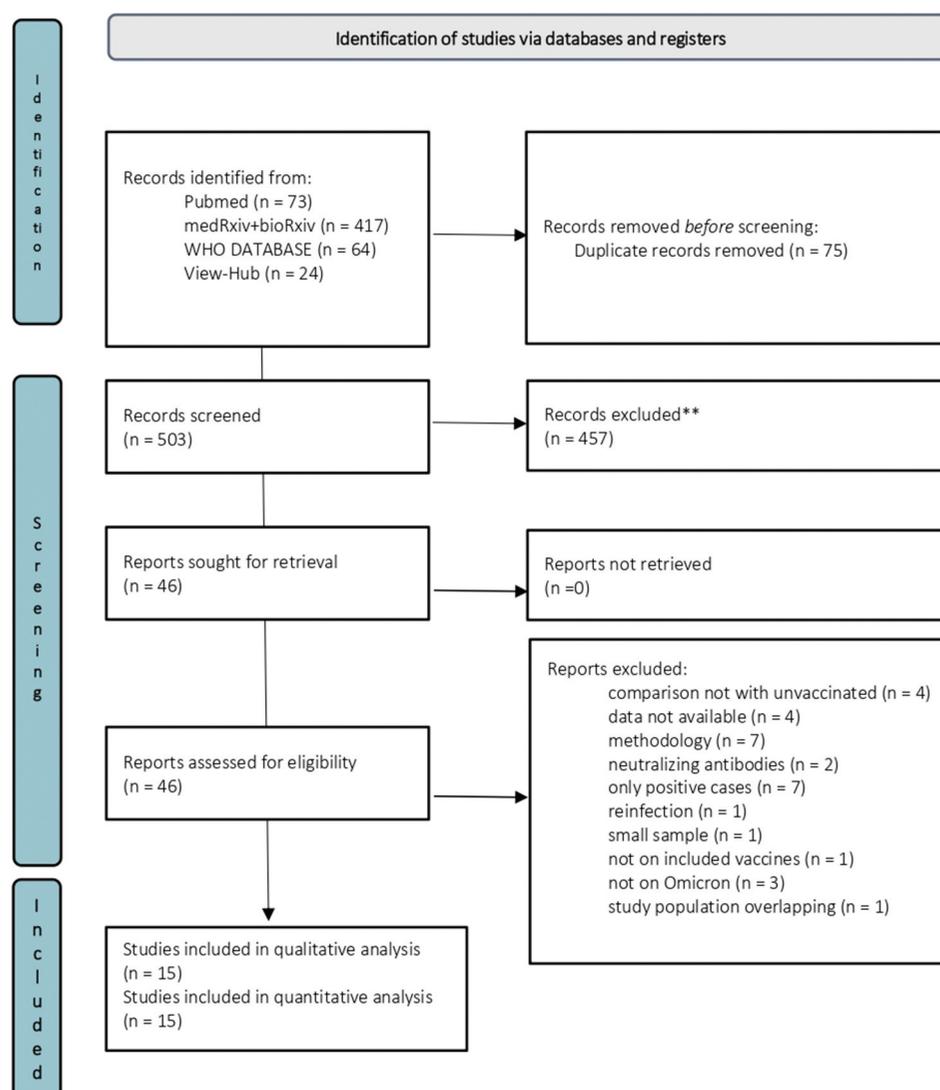


Figure 1. Prisma flow diagram.

### 3.1. Risk of omicron infection after primary course vaccination

A total of 14 studies and  $k = 27$  observations are included in this meta-analysis. The median follow-up period is 213 days (70–365). The observed log odds ratios range from  $-1 \cdot 275$  to  $0 \cdot 467$ , with the majority of estimates being negative (67%). The estimated average log odds ratio based on the RE model is  $\hat{\mu} = -0 \cdot 3788$  (95% CI:  $-0 \cdot 568$  to  $-0 \cdot 190$ ). The values are transformed into the odds ratio scale through exponentiation, such that  $OR = \exp(\hat{\mu}) = 0 \cdot 685$  (95% CI:  $0 \cdot 567$  to  $0 \cdot 827$ ). The average outcome differs significantly from zero ( $z = -3 \cdot 931$ ,  $p < 0 \cdot 0001$ ). Hence, the result suggests that the risk of Sars-Cov2 infection in vaccinated individuals is on average 31·5% lower than the infection risk in unvaccinated. The forest plot is exhibited in. According to the Q-test, the true outcomes appear to be heterogeneous ( $Q(26) = 1962 \cdot 9$ ,  $p < 0 \cdot 0001$ ;  $\tau^2 = 0 \cdot 225$ ;  $I^2 = 99 \cdot 49\%$ ). Neither the rank correlation nor the Egger's regression test indicate any funnel plot asymmetry ( $p = 0 \cdot 901$  and  $p = 0 \cdot 409$ , respectively). The analysis of heterogeneity is displayed in **Figures S1-S5 (Supplementary material)**.

The subgroup analysis includes five subgroups, three of which display significant results ( $p < 0 \cdot 05$ ). Regarding the vaccines used for the primary vaccination, only messenger RNA (mRNA) vaccine exhibits a significant  $OR = 0 \cdot 62$  (95% CI:  $0 \cdot 51$  to  $0 \cdot 76$ ) (**Figure S11, Supplementary material**). The stratified meta-analysis assessing the primary vaccination effectiveness against Omicron VOC by severity of symptoms includes three subgroups and the test for subgroup differences is significant ( $Q_M(df = 2) = 23 \cdot 30$ ,  $p < 0 \cdot 0001$ ) (**Figure 2**). According to the three-level meta-analysis, the 35·6% of the total variance is distributed within the effect sizes (second level), whilst the 64·1% is distributed between groups (third level). The multiple meta-regression embeds four moderators: risk of bias, mean age of the samples (variable centered on the overall mean value of 45 years), vaccine employed in the primary course vaccination (viral vector vaccine or 'VV,' mRNA, and heterologous vaccination with both VV and mRNA or 'VV/mRNA'). Albeit reduced, the residual heterogeneity remains significant ( $QE(df = 19) = 448 \cdot 6$ ,  $p < 0 \cdot 0001$ ;  $\tau^2 = 0 \cdot 0701$ ;  $I^2 = 97 \cdot 25\%$ ). The  $(0.2254 - 0.0701)/0.2254 = 68.9\%$  of the total amount of heterogeneity can be explained by including four moderators in the meta-regression, suggesting further unobserved effects not captured by the model. On average, the risk of symptomatic Covid-19 appears 24% lower for the vaccinated group compared to the unvaccinated ( $OR = 0 \cdot 76$ ; 95% CI:  $0 \cdot 58$  to  $0 \cdot 99$ ), while the risk of hospitalization is 50% lower for the vaccinated group ( $OR = 0 \cdot 50$ ; 95% CI:  $0 \cdot 34$  to  $0 \cdot 72$ ). The OR estimate for any positive rt-PCR is not significant (**Figure 3, Figure S13 and Table S4, Supplementary material**).

### 3.2. Risk of omicron infection after one booster dose

A total of  $k = 28$  observations and 13 studies are included in this meta-analysis. The median follow-up is 62 days (14–150). All the studies investigate the effectiveness of mRNA booster dose except 'Gray,' which demonstrates the efficacy of a two-

dose regimen of Ad26.COV.2 vaccine [32]. The observed log odds ratios range from  $-2 \cdot 5194$  to  $-0 \cdot 0550$ , with the 100% of estimates being negative. The average log odds ratio based on the RE model is  $\hat{\mu} = -1 \cdot 2157$  (95% CI:  $-1 \cdot 4854$  to  $-0 \cdot 9460$ ). Therefore, the outcome differs significantly from zero ( $z = -8 \cdot 8351$ ,  $p < 0 \cdot 0001$ ). The exponentiation yields an average  $OR = 0 \cdot 296$  (95% CI:  $0 \cdot 226$  to  $0 \cdot 388$ ), hence, vaccinated with one booster dose have a 70·4% risk reduction of Omicron infection compared to unvaccinated. According to the Q-test, the true outcomes appear to be heterogeneous ( $Q(27) = 4624 \cdot 51$ ,  $p < 0 \cdot 0001$ ;  $\tau^2 = 0 \cdot 4686$ ;  $I^2 = 99 \cdot 33\%$ ). The influential analysis does not detect overly influential outliers (**Figure S6-S10, Supplementary material**). There is no indication of publication bias because neither the rank correlation nor the regression test indicates any funnel plot asymmetry ( $p = 0 \cdot 7992$  and  $p = 0 \cdot 0735$ , respectively). The subgroup analysis includes six subgroups, four of which display significant results ( $p < 0 \cdot 05$ ). Notably, the risk reduction for the booster group seems 69% lower in studies reporting results at 3 months of follow-up ( $OR = 0 \cdot 31$ ; 95% CI:  $0 \cdot 23$  to  $0 \cdot 42$ ) and 76% in studies reporting 5-months follow-up ( $OR = 0 \cdot 24$ ; 95% CI:  $0 \cdot 12$  to  $0 \cdot 428$ ) at most. However, the test for interaction is not significant (**Figure S12, Supplementary material**).

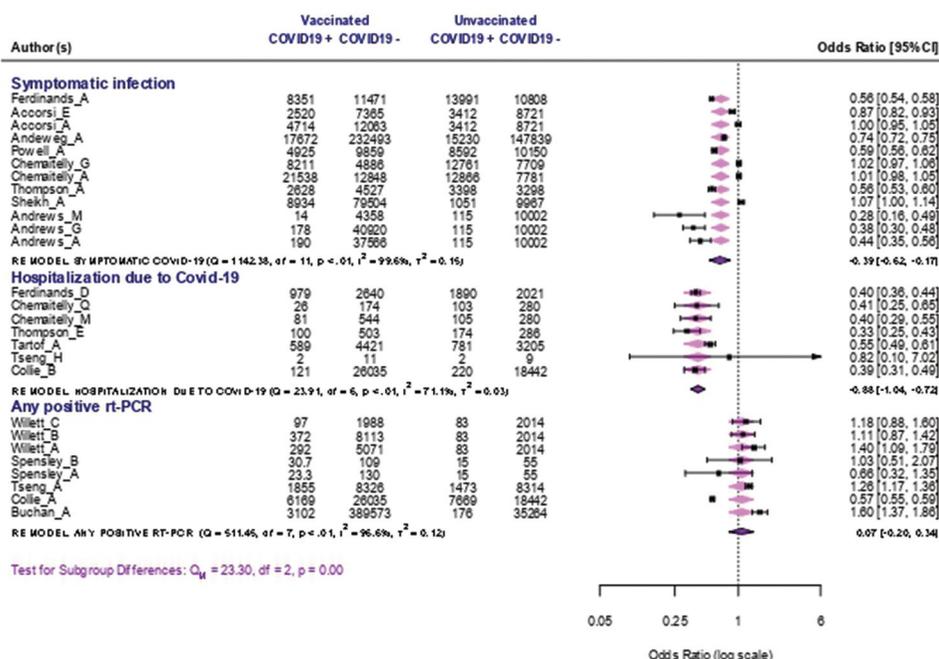
The meta-analysis on one booster effectiveness against Omicron VOC stratified by clinical severity includes three subgroups (**Figure 2b**). The test for subgroup differences is significant ( $QM(df = 2) = 10 \cdot 88$ ,  $p = 0 \cdot 004$ ). According to the multilevel meta-analysis approach, the 57·3% of the total variance is distributed within effect sizes at the second level, whilst the 42·2% is distributed between groups (level 3).

The multiple meta-regression model includes four moderators: risk of bias, the mean age of the samples (centered on the overall mean value of 44·43 years), the regimen of the primary course vaccination ('VV,' mRNA, 'VV/mRNA'). As expected, the residual heterogeneity slightly decreases but remains significant ( $QE(df = 20) = 306 \cdot 9$ ,  $p < 0 \cdot 0001$ ;  $\tau^2 = 0 \cdot 1037$ ;  $I^2 = 94 \cdot 54\%$ ) suggesting further unobserved effect not captured by the predictors in the model. Overall, the multiple meta-regression can explain the 77·9% of the total amount of heterogeneity. On average, the risk of symptomatic Covid-19 appears, 69% lower for the booster group compared to the unvaccinated ( $OR = 0 \cdot 31$ ; 95% CI:  $0 \cdot 23$  to  $0 \cdot 40$ ), whilst the risk of hospitalization is on average 88% lower ( $OR = 0 \cdot 12$ ; 95% CI:  $0 \cdot 08$  to  $0 \cdot 19$ ) (**Figure S14 and Table S4, Supplementary material**).

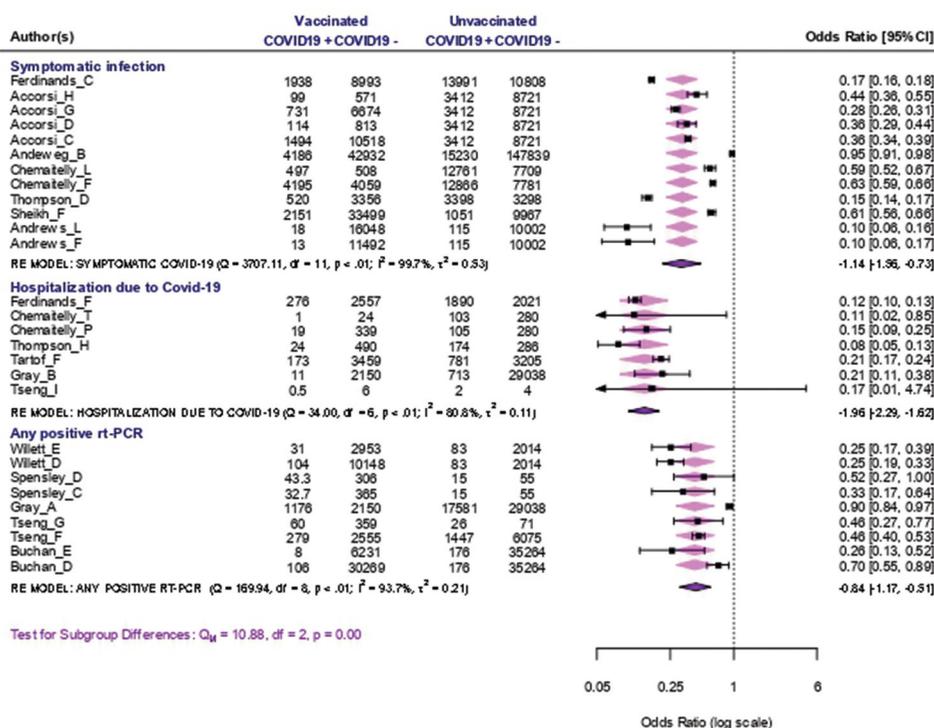
### 3.3. Waning effectiveness of Sars-Cov2 primary vaccination against Omicron VOC

Overall, eight studies assessed the effectiveness of the primary vaccination against Sars-Cov2 at consecutive time intervals. The time intervals correspond to 3 months, 3 to 6 months, 6 months and longer than 6 months since the last dose administration. Therefore, the stratified meta-analyses on Sars-Cov2 vaccine waning effectiveness against Omicron include four subgroups (**Figure 4**). The risk of developing symptomatic Covid-19 is investigated by seven studies and the risk of hospitalization is investigated by four studies (**Table S5,**

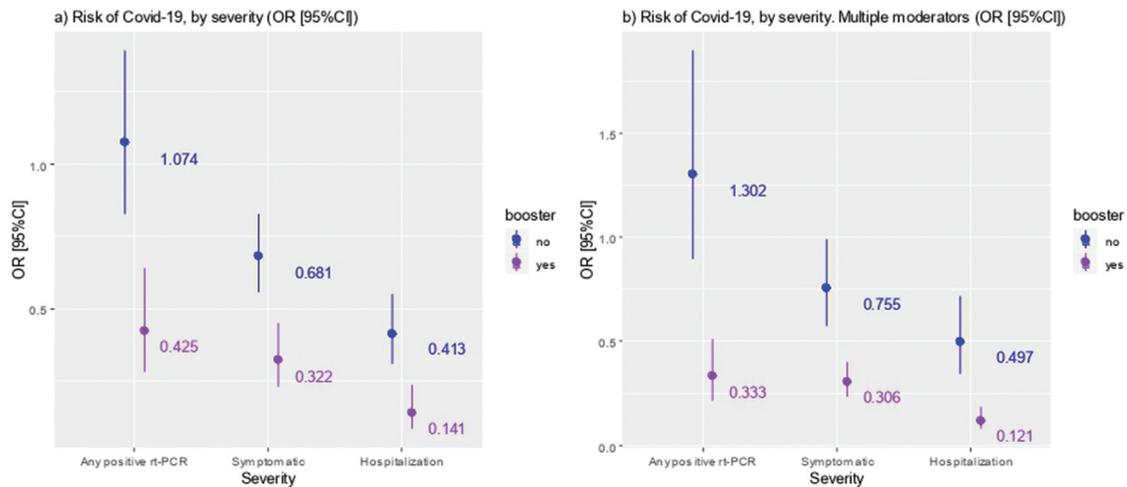
**a) Stratified forest plot: Covid-19 risk by severity of infection after primary vaccination**



**b) Stratified forest plot: Covid-19 risk by severity of infection after one booster**



**Figure 2.** Stratified forest plots and subgroup meta-analyses. Random effect model, IV method. (a) Effectiveness of primary course vaccination, by severity of symptoms. The risk of symptomatic Covid-19 is assessed by 12 observations, the risk of hospitalization by seven, and the risk of any positive rt-PCR by eight. The test for subgroup difference is significant (QM (df = 2) = 23.3, p < 0.0001). According to the subgroup analysis, the risk of any positive rt-PCR appears 7% higher among the vaccinated group with respect to the unvaccinated, however, the result is not significant (OR = 1.07; 95%CI: 0.82 to 1.40). The risk reduction for symptomatic Covid-19 is 32% lower among the vaccinated group compared to the unvaccinated (OR = 0.68; 95%CI: 0.54 to 0.85). Regarding hospitalization due to Omicron infection, the risk appears 58% lower for the vaccinated group compared to the unvaccinated (OR = 0.42; 95%CI: 0.35 to 0.49). (b) Effectiveness of one booster dose against Omicron VOC, by severity of symptoms. The effectiveness of one booster dose is estimated by 12 observations for symptomatic Covid-19, by seven for hospitalization risk, and by nine for any positive rt-PCR. The test for subgroup differences is significant (QM (df = 2) = 10.88, p < 0.0001). The risk of positive rt-PCR appears 57% lower among the booster group with respect to the unvaccinated group (OR = 0.43; 95%CI: 0.31 to 0.60). The risk reduction in favor of the booster group is 68% for symptomatic Covid-19 (OR = 0.32; 95%CI: 0.21 to 0.48) and 86% for hospitalization (OR = 0.14; 95%CI: 0.10 to 0.20).



**Figure 3.** Meta-regression model estimates, risk of Omicron infection by severity of symptoms: OR (95%CI) estimates from meta-regression with one moderator (a) and multiple moderator (b). In the restricted meta-regression (one moderator), the risk of any positive rt-PCR appears 7% higher among the primary vaccination group with respect to unvaccinated ( $OR = 1.07$ ; 95%CI: 0.83 to 1.39); whilst in the multiple meta-regression, the risk of any positive rt-PCR appears nearly 30% higher for primary vaccination. However, the results are not significant ( $OR = 1.302$ ; 95%CI: 0.894 to 1.898).

**Supplementary material).** Only in 33% of cases, Omicron rt-PCR positivity is tested routinely; therefore, it is not possible to consistently estimate the vaccine effectiveness in preventing Sars-Cov2 infection, as well as the vaccine's capability of limiting the virus spreading.

Concerning the risk of symptomatic Omicron infection after vaccination with primary course, a total of  $k = 29$  observations are included in the meta-analysis; all estimates are based on the RE model. The overall Log odds ratio based on the RE model is  $\hat{\mu} = -0.4792$  (95% CI:  $-0.6418$  to  $-0.3165$ ), equivalent to  $OR = 0.62$  (95% CI = 0.53–0.73) after exponentiation. The outcomes appear heterogeneous ( $Q(28) = 1394.37$ ,  $p < 0.0001$ ;  $\tau^2 = 0.1773$ ;  $I^2 = 99.01\%$ ) and the regression test indicates funnel plot asymmetry ( $p < 0.0001$ ); however, the rank correlation test is not significant ( $p = 0.3051$ ) (**Figure S17a, Supplementary material**). The test for subgroup differences is not statistically significant ( $Q_M(df = 3) = 4.169$ ,  $p = 0.2438$ ) (**Figure 4a**). According to the multilevel meta-analysis, the 93.1% of the total variance is distributed at second level ( $\sigma^2 = 0.168$ ), while the 5.9% is distributed at the third level (between groups).

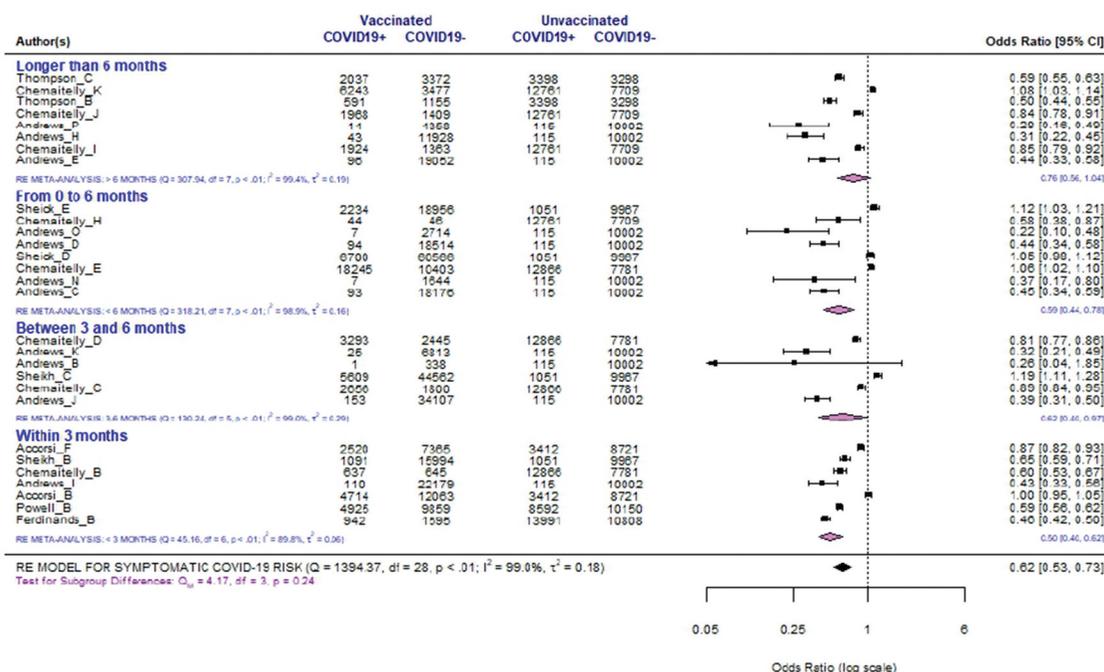
The multiple meta-regression model includes four moderators: time-lapse since the last dose, risk of bias, age of the study (variable centered on the mean value of 41.4), vaccine technology (VV, mRNA, heterologous vaccination, or VV/mRNA). The residual heterogeneity notably decreases but remains significant ( $QE(df = 20) = 53.6$ ,  $p < 0.0001$ ;  $\tau^2 = 0.0042$ ;  $I^2 = 65.10\%$ ). The average risk reduction is 46% for vaccinated with respect to unvaccinated ( $OR = 0.54$ ; 95% CI: 0.48 to 0.61) within 3 months, 22% within 6 months ( $OR = 0.78$ ; 95% CI: 0.69 to 0.88) and 16% between 3 and 6 months ( $OR = 0.84$ ; 95% CI: 0.74 to 0.96). Moreover, the OR decreases on average by 2% in studies where the mean age is one more unit away from the overall mean of 41.4 years ( $OR = 0.98$ ; 95% CI: 0.98 to 0.99). The heterologous vaccination (VV/mRNA) provides a positive coefficient and an 18% higher risk of symptomatic Omicron

infection with respect to mRNA vaccine regimens ( $OR = 1.18$ ; 95% CI: 1.08 to 1.29) (**Figure S20 and Table S6, Supplementary material**).

The meta-analysis on Sars-Cov2 primary vaccination effectiveness against hospitalization embeds a total of 4 studies and 11 observations (**Figure 4b**). The average log odds ratio based on the RE model is  $\hat{\mu} = -0.8634$  (95% CI:  $-1.0348$  to  $-0.6920$ ), which corresponds to  $OR = 0.42$  (95% CI: 0.36 to 0.50) by exponentiation. According to the Q-test, the true outcomes appear heterogeneous ( $Q(10) = 36.38$ ,  $p < 0.0001$ ;  $\tau^2 = 0.051$ ;  $I^2 = 73.3\%$ ). The regression test indicates a funnel plot asymmetry ( $p = 0.0272$ ); however, it is not confirmed by the rank correlation test ( $p = 0.542$ ) (**Figure S17, Supplementary material**). The test for subgroup differences suggests that there is not a statistically significant subgroup effect ( $Q_M(df = 3) = 3.9437$ ,  $p = 0.268$ ). The three-level meta-analysis approach shows that the 73.3% of the total variance is distributed at the second level ( $\sigma^2 = 0.051$ ) while  $\sigma^2 = 0.00$  at third level.

In the multiple meta-regression model, only three predictors are designated as moderators because 'vaccine regimen' contains only observations on mRNA vaccines. The estimated amount of residual heterogeneity is  $\tau^2 = 0.00$  and the test for residual heterogeneity is no longer significant ( $QE(df = 4) = 1.83$ ,  $p = 0.766$ ). All moderators exhibit significant coefficients except 'risk of bias.' The adjusted average effect corresponds to an  $OR = 0.28$  (95% CI = 0.21 to 0.38) within 3 months and average risk reduction of 72% for vaccinated in comparison to unvaccinated. The average OR raises to 0.38 (95% CI: 0.25 to 0.59) within 6 months and to 0.45 (95% CI: 0.30 to 0.68) after more than 6 months. The variable 'age' (centered on the mean of 48.3 years) generates a significant coefficient indicating that the risk of hospitalization increases on average by 2.6% by increasing of one unit the study mean age ( $OR = 1.026$ ; 95% CI: 1.003 to 1.049) (**Figure S21 and Table S6, Supplementary material**). The predicted ORs for symptomatic infection and hospitalization risks are plotted in **Figure 5**.

a) Stratified forest plot: effectiveness of primary vaccination against symptomatic Covid-19, by time intervals



b) Stratified forest plot: effectiveness of primary vaccination against hospitalization risk, by time intervals

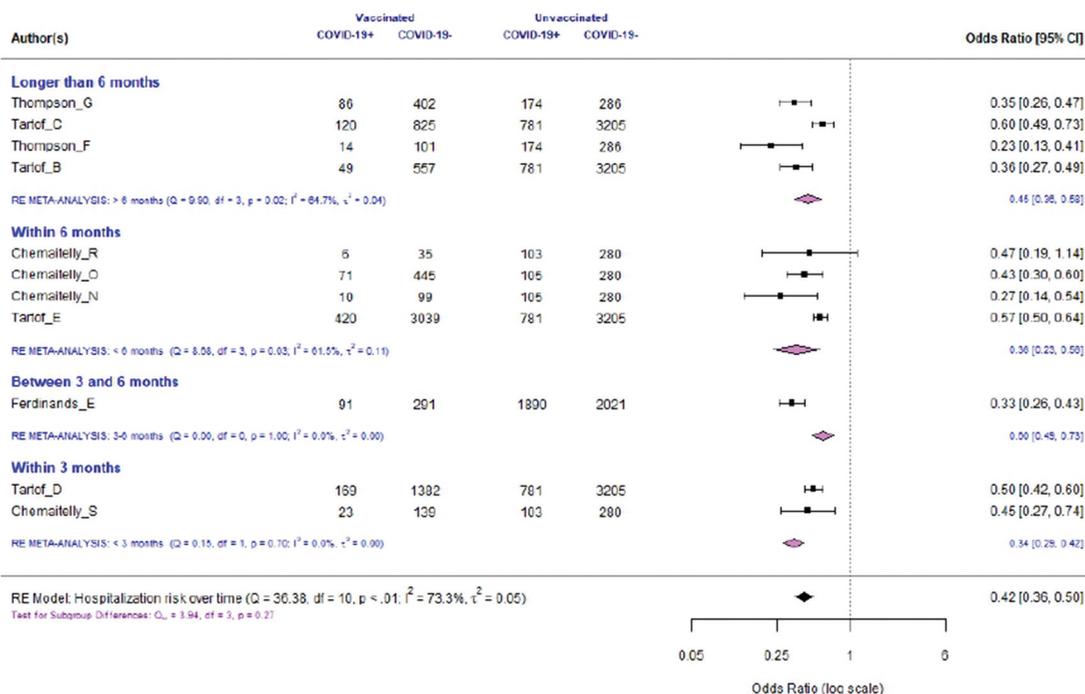
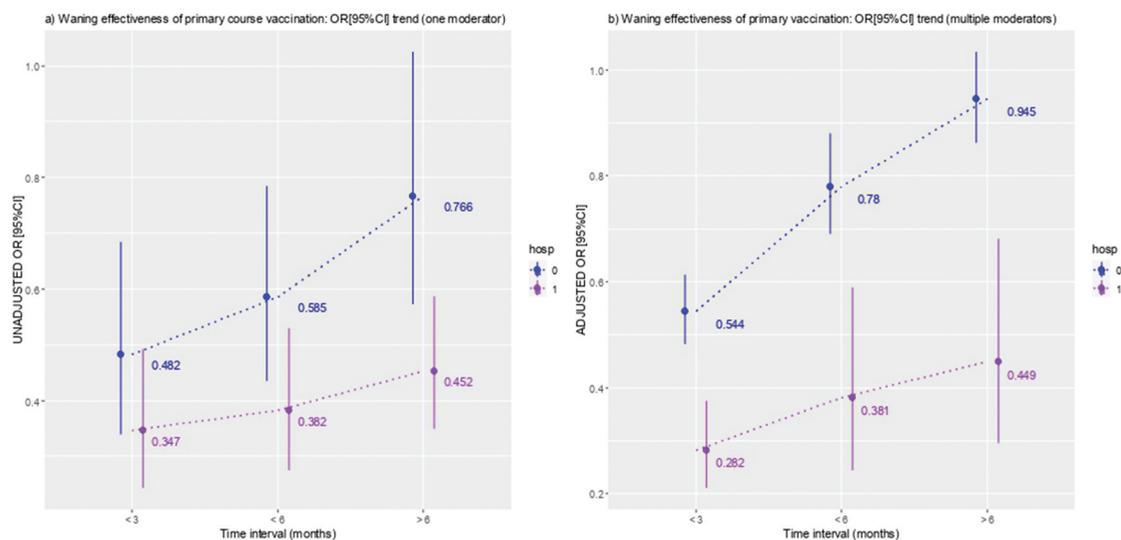


Figure 4. Stratified forest plots. The forest plots include four subgroups representing four discrete time intervals. The results of the individual studies are grouped together according to the corresponding subgroup. Below each subgroup, a summary polygon shows the results of a RE meta-analysis. The pooled effect sizes are expressed as log Odds Ratios. The summary polygon at the bottom of the plot shows the results from the overall RE model (IV method). (a) Stratified forest plot, symptomatic Omicron infection risk, by time intervals. According to the subgroup analysis, the risk reduction appears to be 50% among vaccinated compared to unvaccinated until 3 months (OR = 0.50; 95% CI: 0.40 to 0.62). The risk reduction decreases to nearly 41% with respect to unvaccinated within 6 months (OR = 0.59; 95% CI: 0.44–0.78), and to 24% thereafter (OR = 0.76; 95% CI: 0.56 to 1.03). (b) Stratified forest plot, hospitalization due to Omicron infection risk, by time intervals. The risk reduction appears 65% lower among vaccinated compared to unvaccinated within 3 months (OR = 0.35; 95% CI: 0.29 to 0.42); whereas the overall risk reduction is on average 64% compared to unvaccinated within 6 months (OR = 0.36; 95% CI: 0.23 to 0.56) and 54% thereafter (OR = 0.46; 95% CI = 0.36 to 0.58). Only one study assesses the risk of hospitalization between 3 and 6 months (OR = 0.60; 95% CI = 0.49 to 0.73) (41).



**Figure 5.** Plots displaying the trend of symptomatic Omicron infection and related hospitalization risk. Y-axes: ORs [95%CI] estimates from meta-regression with one moderator (a) and meta-regression with multiple moderators (b). The risk of symptomatic Covid-19 infections is depicted in blue while the risk of hospitalization is in purple. X-axes: time intervals at 3 months, 6 months and over 6 months from last dose administration. Time interval running from 3 to 6 months is suppressed because only one study estimate is available for hospitalization risk.

#### 4. Discussion

The evidence achieved through the quantitative synthesis suggests that a primary vaccination course is not sufficiently protective against Omicron. In fact, the probabilities of symptomatic infection and related hospitalization are nearly 50% for vaccinated with respect to unvaccinated, based on a maximum follow-up of one year. One additional booster dose decreases by 69% the risk of symptomatic Omicron infection (OR = 0.31; 95% CI: 0.23 to 0.40) and by 88% the risk of hospitalization (OR = 0.12; 95% CI: 0.08 to 0.19) with respect to unvaccinated at a maximum follow-up of 5 months. Albeit not significant, the subgroup analysis does not suggest a waning effectiveness of the booster dose after 5 months, however, the evidence on long-term effectiveness is still limited.

The risk of any positive rt-PCR appears higher among the primary vaccination group with respect to the unvaccinated (OR = 1.302; 95% CI: 0.89 to 1.90); however, the results are not significant.

Age does not appear as a significant predictor, notwithstanding the negative association with the overall risk of infection after the primary vaccination and after the booster. Conversely, age is negatively associated with the risk of symptomatic infection and positively associated with the risk of hospitalization after the primary vaccination. Some unobserved effect of uncontrolled confounding must be acknowledged in interpreting this association. For instance, the different extent to which the joint effect of the mitigation measures uplift has affected the younger and the elderly population. However, despite the generalizability, these results do not allow us to infer any clear conclusion.

There is no clear advantage between homologous and heterologous vaccination, particularly on boosting, probably because the majority of the appraisals have been conducted on mRNA vaccination and data on heterologous vaccination are quite sparse.

As the administration of booster doses, whether homologous or heterologous, should take into consideration the waning protection of the primary course and the optimal interval for an efficient immune response, the implications of our findings extend to health care and public health policy.

Our results on the waning trends align with the estimates provided by the clinical trials [43–45]. According to our estimates, the effectiveness of primary vaccination against Omicron reaches a peak within 3 months determining a risk reduction of roughly 72% with respect to unvaccinated. The protection is maintained at 6 months, with a risk reduction of nearly 62%, and dramatically declines thereafter (55% less probability for vaccinated compared to unvaccinated). Overall, the effectiveness against hospitalization diminishes by approximately 10–15% every 3 months, and the point estimates show wide confidence intervals [46].

The ramping-up trend for symptomatic Omicron infection risk appears steeper than the trend for hospitalization risk; in other words, the protection against symptomatic Covid-19 declines faster. The risk reduction of symptomatic Omicron infection after a primary vaccination declines sharply to 22% in 6 months.

Our study provides the best available data synthesis on vaccine effectiveness against Omicron; however, several limitations must be acknowledged. First, only in 33% of cases, Omicron rt-PCR positivity is tested routinely; therefore, it is not possible to draw conclusions about vaccine effectiveness in preventing Omicron infection. Second, by examining periods during which Omicron and Delta coexistence was very likely, early studies generate a distortion of the VE effectiveness estimate.

In part, the high heterogeneity surrounding the meta-analysis estimates stems from the observational design of the included studies. Unless a randomization process, the meta-regression cannot capture the unobserved effect of

confounders such as the level of community transmission, the implementation of public health prevention measures, and the spread of new variants. For instance, regarding the Omicron variant definition, the studies on BA.1 do not distinguish between the different sub-lineages, although the majority of them are conducted during the BA.1 surge. Differences between BA.2 and BA.1 in evading immunity remain undefined.

## 5. Conclusion

In conclusion, despite the high heterogeneity, only in part explained by the meta-regressions, this study confirms that primary vaccination does not provide sufficient protection against symptomatic Omicron infection, because the overall estimate of effectiveness never reaches a minimum requirement of 50% in the risk reduction. One additional booster dose decreases substantially the risk of symptomatic Omicron infection and of hospitalization. The booster-dose administration should be recommended after 3 months and no later than 6 months following the primary vaccination course.

## 6. Expert opinion

The findings of this systematic review and meta-analysis provide further knowledge about the effectiveness of the primary vaccination and the administration of one additional booster dose, against different outcomes, such as infection, symptomatic disease, and hospitalization. Despite the high heterogeneity, only in part explained by the meta-regressions, this study confirms that primary vaccination does not provide sufficient protection against symptomatic Omicron infection, because the overall effectiveness estimate never reaches a minimum requirement of 50% in risk reduction. One additional booster dose decreases substantially the risk of symptomatic Omicron infection and of hospitalization. The booster dose administration should be recommended after 3 months and no later than 6 months following the primary vaccination course. Real-world data provide a clearer picture of this pandemic dynamics, and consequently of clinical outcomes, compared to neutralization and modeling studies. To some extent, public health recommendations on the choice and timing of vaccine schedules should be driven by real-world studies, even though sometimes the results are produced too late. In our case, the review is based on available studies, which in most cases have been conducted in the early Omicron era. Nonetheless, they represent the best available real-world evidence on the effect of current vaccines against this new variant.

Any systematic review suffers from methodological limitations originated by primary studies, i.e. short study duration, study design or study population heterogeneity, that could be addressed by sharing/planning a comprehensive protocol among different study centers, possibly following an ongoing-updated living methodology on surveillance-clinical data. However, we do think that our analysis may help to better quantify the effectiveness of currently available vaccines against an emerging variant characterized by immune evasion.

## Funding

This paper was not funded.

## Declaration of interest

The authors have no relevant affiliations or financial involvement with any organization or entity with a financial interest in or financial conflict with the subject matter or material discussed in the manuscript. This includes employment, consultancies, honoraria, stock ownership or options, expert testimony, grants or patents received or pending, or royalties.

## Reviewer disclosures

Peer reviewers on this manuscript have no relevant financial or other relationships to disclose.

## Data availability

Data supporting the reported results are available on request to the Authors.

## References

1. GISAID - hCov19 Variants . [Accessed 7 June 2022 . : <https://www.gisaid.org/hcov19-variants/>
2. Peacock TP, Jc B, Zhou J, et al. The altered entry pathway and antigenic distance of 2 the SARS-CoV-2 Omicron variant map to separate 3 domains of spike protein bioRxiv. [cited 2021 Jun 7]. 2022. DOI: [10.1101/2021.12.31.474653](https://doi.org/10.1101/2021.12.31.474653)
3. Abdelnabi R, Foo CS, Zhang X, et al. The omicron (B.1.1.529) SARS-CoV-2 variant of concern does not readily infect Syrian hamsters .bioRxiv .cited 2022 Jun 7]; 2021.12.24.474086.Available from: <https://www.biorxiv.org/content/10.1101/2021.12.24.474086v1> 2021 Dec 26
4. Viana R, Moyo S, Amoako DG, et al. Rapid epidemic expansion of the SARS-CoV-2 Omicron variant in Southern Africa. Nature. Internet]. 2022 Mar 24;603(7902):679–686. Available from. <https://pubmed.ncbi.nlm.nih.gov/35042229/>
5. Yang W, Shaman J Dec 21 [revised 2022 Jun 29] . SARS-CoV-2 transmission dynamics in South Africa and epidemiological characteristics of the omicron variant. medRxiv 2021 doi:10.1101/2021.12.19.21268073 07 Jun 2022 . <http://www.ncbi.nlm.nih.gov/pubmed/34981071>
6. [PDF] Report 50: hospitalisation risk for Omicron cases in England | semantic scholar. [cited 2022 Jun 7]. Available from: <https://www.semanticscholar.org/paper/Report-50%3A-Hospitalisation-risk-for-Omicron-cases-Ferguson-Ghani/8b696e1a05092a11b39b92b4dea7bc05e7032df>
7. Bager P, Wohlfahrt J, Bhatt S, et al. Reduced risk of hospitalisation associated with infection with SARS-CoV-2 Omicron relative to delta: a Danish cohort study SSRN Electron J 2022Jan20 cited 2022 Jun 7. <https://papers.ssrn.com/abstract=4008930>
8. Lewnard JA, Hong VX, Patel MM, et al. Clinical outcomes associated with Omicron (B.1.1.529) variant and BA.1/BA.1.1 or BA.2 subvariant infection in southern California medRxiv . [cited 2022 Jun 7]; cited: [10.1101/2022.01.11.22269045](https://doi.org/10.1101/2022.01.11.22269045)
9. Nyberg T, Ferguson NM, Nash SG, et al. Comparative analysis of the risks of hospitalisation and death associated with SARS-CoV-2 omicron (B.1.1.529) and delta (B.1.617.2) variants in England: a cohort study. Lancet. 2022 Apr 2 [cited 2022 Jun 7];399(10332):1303–1312. Available from: <http://www.thelancet.com/article/S0140673622004627/fulltext>
10. Hayawi K, Shahriar S, Serhani MA, et al. Vaccine versus variants (3Vs): are the COVID-19 vaccines effective against the variants? A systematic review. Vaccines (Basel). 2021 Nov 1;9(11). [10.3390/vaccines9111305](https://doi.org/10.3390/vaccines9111305)

11. Harder T, Külper-Schiek W, Reda S, et al. Effectiveness of COVID-19 vaccines against SARS-CoV-2 infection with the Delta (B.1.617.2) variant: second interim results of a living systematic review and meta-analysis, 1 January to 25 August 2021. *Euro Surveill.* 2021 Oct 14;26(41). [cited 2022 Jun 7]; Available from: <https://pubmed.ncbi.nlm.nih.gov/34651577/>
12. Liu Q, Qin C, Liu M, et al. Effectiveness and safety of SARS-CoV-2 vaccine in real-world studies: a systematic review and meta-analysis. *Infect Dis Poverty* [Internet]. 2021 Dec 1 cited 2022 Jun 7];10(1). Available from: [/pmc/articles/PMC8590867/ 10.1186/s40249-021-00915-3](https://pubmed.ncbi.nlm.nih.gov/340249-021-00915-3)
13. Kow CS, Ramachandram DS, Hasan SS. The effectiveness of mRNA-1273 vaccine against COVID-19 caused by Delta variant: a systematic review and meta-analysis. *J Med Virol.* 2022 May 1 [cited 2022 Jun 7];94(5):2269–2274. Available from: <https://pubmed.ncbi.nlm.nih.gov/34978339/>
14. Pormohammad A, Zarei M, Ghorbani S, et al. Effectiveness of COVID-19 vaccines against delta (B.1.617.2) variant: a Systematic review and meta-analysis of clinical studies. *Vaccines (Basel)*. Internet]. 2021 Jan 1 [cited 2022 Jun 7];10(1). Available from ;(). <https://pubmed.ncbi.nlm.nih.gov/35062684/>
15. Cele S, Jackson L, Khoury DS, et al. SARS-CoV-2 Omicron has extensive but incomplete escape of Pfizer BNT162b2 elicited neutralization and requires ACE2 for infection *medRxiv* 2021Dec17 cited 2022 Jun 7. <http://www.ncbi.nlm.nih.gov/pubmed/34909788>
16. Wilhelm A, Widera M, Grikscheit K, et al. Reduced neutralization of SARS-CoV-2 Omicron variant by vaccine sera and monoclonal antibodies. *medRxiv* [Internet]. 2021 Dec 8 [cited 2022 Jun 7]; Available from: <https://www.medrxiv.org/content/10.1101/2021.12.07.21267432v1>
17. Schmidt F, Muecksch F, Weisblum Y, et al. Plasma neutralization properties of the SARS-CoV-2 Omicron variant. *medRxiv.* 2021 Dec 13 cited 2022 Jun 7 2021.12.12.21267646. <https://www.medrxiv.org/content/10.1101/2021.12.12.21267646v1>
18. Barouch DH, Stephenson KE, Sadoff J, et al. Durable Humoral and cellular immune responses following Ad26.COV2.S vaccination for COVID-19. *medRxiv.* 2021 Jul 7 cited 2022 Jun 8 2021.07.05.21259918. <https://www.medrxiv.org/content/10.1101/2021.07.05.21259918v1>
19. Doria-Rose N, Suthar MS, Makowski M, et al. Antibody persistence through 6 months after the second dose of mRNA-1273 Vaccine for Covid-19. *N Engl J Med.* 2021 Jun 10 [cited 2022 Jun 7];384(23):2259–2261. Available from: <https://pubmed.ncbi.nlm.nih.gov/33822494/>
20. Naaber P, Tserel L, Kangro K, et al. Dynamics of antibody response to BNT162b2 vaccine after six months: a longitudinal prospective study *The Lancet regional health Europe.* 2021 Nov 1 [cited 2022 Jun 7];10. . <https://pubmed.ncbi.nlm.nih.gov/34514454/>
21. Garcia-Beltran WF, St. Denis KJ, Hoelzemer A, et al. mRNA-based COVID-19 vaccine boosters induce neutralizing immunity against SARS-CoV-2 Omicron variant. *Cell.* 2022 Feb 3 [cited 2022 Jun 7];185(3):457–466.e4. Available from: <https://pubmed.ncbi.nlm.nih.gov/34995482/>
22. Eroglu B, Nuwarda RF, Ramzan I, et al. Review of COVID-19 vaccines. *Vaccines (Basel)*. Internet]. 2021 Jan 1 [cited 2022 Jun 7];10(1). Available from () <https://pubmed.ncbi.nlm.nih.gov/35062723/>.
23. Page MJ, McKenzie JE, Bossuyt PM, et al. The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. *BMJ*. Internet]. 2021 Mar 29 [cited 2022 Jun 7];372. Available from; <https://pubmed.ncbi.nlm.nih.gov/33782057/>
24. Sterne JA, Hernán MA, Reeves BC, et al. ROBINS-I: a tool for assessing risk of bias in non-randomised studies of interventions. *BMJ.* 2016 Oct 12 cited 2022 Jun 7 355. <https://www.bmj.com/content/355/bmj.i4919>
25. Viechtbauer W. Bias and efficiency of meta-analytic variance estimators in the random-effects model. *J Educ Behav Stat.* 2005;30(3):261–293.
26. Cochran WG. The combination of estimates from different experiments. *Biometrics.* cited 2022 Jun 7];10(1):101.Available from: [/record/1955-00087-001](https://pubmed.ncbi.nlm.nih.gov/record/1955-00087-001).
27. Higgins JPT, Thompson SG. Controlling the risk of spurious findings from meta-regression. *Stat Med.* 2004 Jun 15 [cited 2022 Jun 7];23(11):1663–1682. Available from: <https://pubmed.ncbi.nlm.nih.gov/15160401/>
28. Viechtbauer W, Cheung MWL. Outlier and influence diagnostics for meta-analysis. *Res Synth Methods.* 2010 Apr [cited 2022 Jun 7];1(2):112–125. Available from: <https://pubmed.ncbi.nlm.nih.gov/26061377/>
29. Begg CB, Mazumdar M. Operating characteristics of a rank correlation test for publication bias. *Biometrics.* 1994 Dec;50(4):1088.
30. Sterne JAC, Egger M Regression Methods to Detect Publication and Other Bias in Meta-Analysis Publication Bias in Meta-Analysis: Prevention, Assessment and Adjustments. 2006;Jan.
31. Spensley KJ, Gleeson S, Martin P, et al. Comparison of vaccine effectiveness against the omicron (B.1.1.529) variant in hemodialysis patients. *Kidney Int Rep.* 2022 Apr cited 2022 Jun 7 7: (6) Available from <https://pubmed.ncbi.nlm.nih.gov/35434428/>
32. Gray MBBCH GE, Collie S, Garrett MBBS N, et al. Vaccine effectiveness against hospital admission in South African health care workers who received a homologous booster of Ad26.COV2 during an omicron COVID19 wave: preliminary results of the sisonke 2 study. *medRxiv.* Internet]. 2021 Dec 29 [cited 2022 Jun 7];2021.12.28.21268436. Available from. <https://www.medrxiv.org/content/10.1101/2021.12.28.21268436v1>
33. Tseng HF, Ackerson BK, Luo Y, et al. Effectiveness of mRNA-1273 against SARS-CoV-2 Omicron and delta variants. *Nat Med.* cited 2022 Jun 7];28(5). Available from 2022 May 1;(). <https://pubmed.ncbi.nlm.nih.gov/35189624/>
34. Accorsi EK, Britton A, Fleming-Dutra KE, et al. Association between 3 doses of mRNA COVID-19 vaccine and symptomatic infection caused by the SARS-CoV-2 Omicron and delta variants. *JAMA*. Internet]. 2022 Feb 15 [cited 2022 Jun 7 327(7):639–651. Available from. ;(). <https://pubmed.ncbi.nlm.nih.gov/35060999/>
35. Andeweg SP, Gier B, De, Eggink D, et al. Protection of COVID-19 vaccination and previous infection against omicron BA.1, BA.2 and Delta SARS-CoV-2 infections. *medRxiv.* 2022 May 12 cited 2022 Jun 7 2022.02.06.22270457. <https://www.medrxiv.org/content/10.1101/2022.02.06.22270457v3>
36. Chemaitelly H, Tang P, Hasan MR, et al. Waning of BNT162b2 vaccine protection against SARS-CoV-2 Infection in Qatar. *N Engl J Med.* 2021 Dec 9 [cited 2022 Jun 7];385(24):e83. Available from: <https://www.nejm.org/doi/full/10.1056/NEJMoa2114114>
37. Powell AA, Kirsebom F, Stowe J, et al. Adolescent vaccination with BNT162b2 (Comirnaty, Pfizer-BioNTech) vaccine and effectiveness of the first dose against COVID-19: national test-negative case-control study, England. *medRxiv* [Internet 2021 Dec 11 [cited 2022 Jun 7];2021.12.10.21267408. Available from: <https://www.medrxiv.org/content/10.1101/2021.12.10.21267408v1>
38. Sheikh A, Kerr S, Woolhouse M, et al. Severity of Omicron variant of concern and vaccine effectiveness against symptomatic disease: national cohort with nested test negative design study In: Balint SG, Antala B, Carty C, et al., editors. *Uniwersytet śląski* [Internet]. 2013 [cited 2022 Jun 7];343–354. Available from: <https://sbc.org.pl/dlibra/publication/99008/edition/93276/synteza-aktywnosc-biologiczna-nowych-analogow-tiosemikarbazonowych-chelatorow-zelaza-serda-maciej?language=en>
39. Collie S, Champion J, Moultrie H, et al. Effectiveness of BNT162b2 vaccine against omicron variant in South Africa. *N Engl J Med.* 2022 Feb 3 [cited 2022 Jun 8];386(5):494–496. Available from: <https://www.nejm.org/doi/full/10.1056/NEJM2119270>
40. Ferdinands JM, Rao S, Dixon BE, et al. Waning 2-Dose and 3-dose effectiveness of mRNA vaccines against COVID-19-associated emergency department and urgent care encounters and hospitalizations among adults during periods of delta and omicron variant predominance — VISION network, 10 States, August 2021–January 2022. *MMWR Morb Mortal Wkly Rep.* 2022 cited 2022 Jun

- 8 71: (7) 255–263. Available from <https://www.cdc.gov/mmwr/volumes/71/wr/mm7107e2.htm>
41. Tartof SY, Slezak JM, Puzniak L, et al. BNT162b2 (Pfizer–Biontech) mRNA COVID-19 vaccine against omicron-related hospital and emergency department admission in a large us health system: a test-negative design SSRN Electron J 2022Jan20 cited 2022 Jun 8. <https://papers.ssrn.com/abstract=4011905>
42. Thompson MG, Natarajan K, Irving SA, et al. Effectiveness of a third dose of mRNA vaccines against COVID-19–associated emergency department and urgent care encounters and hospitalizations among adults during periods of delta and omicron variant predominance — VISION network, 10 States, August 2021–January 2022. MMWR Morb Mortal Wkly Rep. 2022 Jan 21 [cited 2022 Jun 8];71(4):139–145. Available from: <https://www.cdc.gov/mmwr/volumes/71/wr/mm7104e3.htm>
43. Polack FP, Thomas SJ, Kitchin N, et al. Safety and Efficacy of the BNT162b2 mRNA Covid-19 Vaccine. N Engl J Med . 2020 Dec 31 [cited 2022 Jun 8];383(27):2603–2615. Available from: <https://pubmed.ncbi.nlm.nih.gov/33301246/>
44. Baden LR, El Sahly HM, Essink B, et al. Efficacy and Safety of the mRNA-1273 SARS-CoV-2 Vaccine. N Engl J Med. 2021 Feb 4;384(5):403–416.
45. Sadoff J, Gray G, Vandebosch A, et al. Safety and Efficacy of Single-Dose Ad26.COV2.S vaccine against Covid-19. N Engl J Med. 2021 Jun 10 [cited 2022 Jun 8];384(23):2187–2201. Available from: <https://www.nejm.org/doi/full/10.1056/NEJMoa2101544>
46. Lin DY, Gu Y, Wheeler B, et al. Effectiveness of Covid-19 vaccines in the United States over 9 months: surveillance data from the state of North Carolina. medRxiv . 2021 Oct 26 cited 2022 Jun 8 2021.10.25.21265304. <https://www.medrxiv.org/content/10.1101/2021.10.25.21265304v1>